

# 1 Hypothesis Testing - Uniformly Most Powerful Tests

We give the definition of a uniformly most powerful test in a general setting which includes one-sided and two-sided tests. We take the null hypothesis to be

$$H_0 : \theta \in \Omega_0$$

and the alternative to be

$$H_1 : \theta \in \Omega_1$$

We write the power function as  $Pow(\theta, d)$  to make its dependence on the decision function explicit.

*Definition:* A decision function  $d^*$  is a uniformly most powerful (UMP) decision function (or test) at significance level  $\alpha_0$  if

- (1)  $Pow(\theta, d^*) \leq \alpha_0, \quad \forall \theta \in \Omega_0$
- (2) For every decision function  $d$  which satisfies (1), we have  $Pow(\theta, d) \leq Pow(\theta, d^*), \quad \forall \theta \in \Omega_1$ .

Do UMP tests ever exist? If the alternative hypothesis is one-sided then they do for certain distributions and statistics. We proceed by defining the needed property on the population distribution and the statistic.

*Definition:* Let  $T = t(X_1, X_2, \dots, X_n)$  be a statistic. Let  $f(x_1, x_2, \dots, x_n | \theta)$  be the joint density of the random sample. We say that  $f(x_1, x_2, \dots, x_n | \theta)$  has a *monotone likelihood ratio in the statistic  $T$*  if for all  $\theta_1 < \theta_2$  the ratio

$$\frac{f(x_1, \dots, x_n | \theta_2)}{f(x_1, \dots, x_n | \theta_1)}$$

depends on  $x_1, \dots, x_n$  only through  $t(x_1, \dots, x_n)$  and the ratio is an increasing function of  $t(x_1, \dots, x_n)$ .

**Example:** Consider a Bernoulli distribution for the population, i.e., we are looking at a population proportion. So each  $X_i = 0, 1$  and  $p = P(X_i = 1)$ . The joint density is

$$f(x_1, \dots, x_n | p) = p^{n\bar{x}}(1 - p)^{n - n\bar{x}}$$

where

$$\bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$$

Let  $p_1 < p_2$ . We have

$$\frac{f(x_1, \dots, x_n | p_2)}{f(x_1, \dots, x_n | p_1)} = \left[ \frac{p_2(1-p_1)}{p_1(1-p_2)} \right]^{n\bar{x}} \left[ \frac{1-p_2}{1-p_1} \right]^n$$

So the ratio depends on the sample only through the sample mean  $\bar{x}$  and it is an increasing function of  $\bar{x}$ . (It is an easy algebra exercise to check that if  $p_2 > p_1$  then  $p_2(1-p_1)/(p_1(1-p_2)) > 1$ .)

**Example:** Now consider a normal population with unknown mean  $\mu$  and known variance  $\sigma^2$ . So the joint density is

$$f(x_1, \dots, x_n | \mu) = \frac{1}{(2\pi)^{n/2} \sigma^n} \exp\left(-\frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \mu)^2\right)$$

Now let  $\mu_1 < \mu_2$ . A little algebra shows

$$\frac{f(x_1, \dots, x_n | \mu_2)}{f(x_1, \dots, x_n | \mu_1)} = \exp\left(\frac{n}{\sigma^2} \bar{x}(\mu_2 - \mu_1) + \frac{(\mu_1^2 - \mu_2^2)n}{2\sigma^2}\right)$$

So the ratio depends on  $x_1, x_2, \dots, x_n$  only through  $\bar{x}$ , and the ratio is an increasing function of  $\bar{x}$ .

**Theorem 1.** Suppose  $f(x_1, \dots, x_n | \theta)$  has a monotone likelihood ratio in the statistic  $T = t(X_1, \dots, X_n)$ . Consider hypothesis testing with alternative hypothesis  $H_a : \theta > \theta_1$ , and null hypothesis  $H_0 : \theta \leq \theta_0$  or  $H_0 : \theta = \theta_0$ . Let  $\alpha_0, c$  be constants such that  $P(T \geq c) = \alpha_0$ . Then the test that rejects the null hypothesis if  $T \geq c$  is a UMP test at significance level  $\alpha_0$ .

**Example:** We continue the example of a normal population with known variance and unknown mean. We saw that the likelihood ratio is monotone in the sample mean. So if we reject the null hypothesis when  $\bar{X}_n \geq c$ , this will be a UMP test with significance level  $\alpha = P(\bar{X}_n \geq c | \mu_0)$ . Given a desired significance level  $\alpha$ , we choose  $c$  so this equation holds. Then the theorem

tells us we have a UMP test. So for every  $\mu > \mu_0$ , our test makes  $Pow(\mu)$  as large as possible.

**Example:** We continue the example of a Bernoulli distribution for the population (population proportion). To be concrete, suppose the null hypothesis is  $p \leq 0.1$  and the alternative is  $p > 0.1$ . We have a random sample of size  $n = 20$ . Let  $\bar{X}$  be the sample proportion. By what we've already done, the test that reject the null hypothesis when  $\bar{X} \geq c$  will be a UMP test. We want to choose  $c$  so that  $P(\bar{X} \geq c) = \alpha_0$ . However,  $\bar{X}$  is a discrete RV (it can only be  $0/20, 1/20, 2/20, \dots, 19/20, 1$ ), so this is not possible. Suppose we want a significance level of 0.005 Using your favorite software (or a table of the binomial distribution) we find that  $P(\bar{x} \geq 6/20 | p = 0.1) = 0.0113$  and  $P(\bar{x} \geq 7/20 | p = 0.1) = 0.0024$ . So we must take  $c = 7/20$ . Then the test that rejects the null if  $\bar{x} \geq 7/20$  is a UMP test at significance level 0.005.

What about two-sided alternatives? It can be shown that there is no UMP test in this setting.